

A NOTE ON THE ESTIMATION OF MEAN OF A SYMMETRICAL POPULATION

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SUMMARY

A refinement of Sahai and Ray estimator has been suggested for estimating the mean \bar{Y} of a population with variance σ^2 . It has the same mean square error but smaller bias than the one suggested by Sahai and Ray.

INTRODUCTION

For estimating the population mean, \bar{Y} Sahai and Ray [3] proposed the estimator

$$te^* = \bar{y} + \frac{s^2 \bar{y}}{(n\bar{y}^2 + s^2)} \quad \dots(1)$$

and obtained expressions for bias and mean square error to the order $O(\bar{n}^2)$.

$$\text{Bias}(te^*) = \frac{c\bar{Y}}{n} \quad \dots(2)$$

and

$$M(te^*) = \frac{\sigma^2}{n} \left(1 - \frac{c}{n} \right) \quad \dots(3)$$

\bar{y} is the sample mean, s^2 the sample variance, $c \left(= \frac{\sigma^2}{\bar{Y}^2} \right)$ the square of the coefficient of variation (supposed to be known) and n is the size of the sample. If c is unknown, then it is a common practice to replace c by its sample value $Ce = \frac{s^2}{\bar{y}^2}$.

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Now we consider the estimator

$$te^{**} = \bar{y} + \frac{s^2 \bar{y}}{(n\bar{y}^2 + s^2)} - \frac{s^4 \bar{y}}{(n\bar{y}^2 + s^2)^2} \dots (4)$$

To evaluate the bias and mean square error of the estimator te^{**} to the order $O(\bar{n}^2)$, let

$$\bar{y} = \bar{Y} + U, \quad s^2 = \sigma^2 + V \dots (5)$$

Where U and V are order $O(\bar{n}^{\frac{1}{2}})$ with $E(U) = E(V) = 0$. Then upto the order $O(\bar{n}^2)$, we have

$$\begin{aligned} te^{**} &= \bar{y} + \frac{s^2 \bar{y}}{(n\bar{y}^2 + s^2)} - \frac{s^4 \bar{y}}{(n\bar{y}^2 + s^2)^2} \\ &= \bar{y} + \left(\frac{s^2}{n\bar{y}} \right) \left(1 - \frac{s^2}{n\bar{y}^2} \right) - \frac{s^4 \bar{y}}{(n\bar{y}^2 + s^2)^2} \\ &= (\bar{Y} - U) + \bar{Y} \left\{ \left(c + \frac{V}{\bar{Y}^2} \right) \left(1 - \frac{U}{\bar{Y}} + \frac{U^2}{\bar{Y}^2} \right) \right. \\ &\quad \left. - \frac{c^2}{n} - \frac{c^2 \bar{Y}}{n^2} \right\} - \frac{c^2 \bar{Y}}{n^2} \\ &= \bar{Y} \left(1 + \frac{U}{\bar{Y}} \right) + \bar{Y} \left(\frac{c}{n} + \frac{V}{n\bar{Y}} - \frac{cU}{n\bar{Y}} + \frac{cU^2}{n\bar{Y}^2} - \frac{c^2}{n^2} - \frac{UV}{n\bar{Y}^3} \right. \\ &\quad \left. + \frac{U^2 V}{n\bar{Y}^4} \right) - \frac{c^2 \bar{Y}}{n^2} \end{aligned}$$

So that

$$\begin{aligned} E(te^{**} - \bar{Y}) &= \frac{c\bar{Y}}{n} + \frac{c^2 \bar{Y}}{n^2} - \frac{2c^2 \bar{Y}}{n^2} \\ &= \frac{c\bar{Y}}{n} \left(1 - \frac{c}{n} \right) \\ \therefore \text{Bias} &= \frac{c\bar{Y}}{n} \left(1 - \frac{c}{n} \right) \dots (6) \end{aligned}$$

and

$$M(te^{**}) = \frac{\sigma^2}{n} \left(1 - \frac{c}{n} \right) \dots (7)$$

From equation (2), (3), (6) and (7) it is seen that

$$B(te^{**}) < B(te^*),$$

and

$$M(te^{**}) = M(te^*)$$

Therefore our estimator te^{**} is superior to the estimator te^* in the sense that it has a smaller bias than the estimator [3]. The percentage reduction in the bias by the proposed estimator is given by

$$\begin{aligned} & \frac{B(te^*) - B(te^{**})}{B(te^*)} \times 100 \\ &= \frac{\frac{c\bar{Y}}{n} - \frac{c\bar{Y}}{n} \left(1 - \frac{c}{n}\right)}{\frac{c\bar{Y}}{n}} \times 100 \\ &= \frac{c}{n} \times 100 \quad \dots (8) \end{aligned}$$

From (8), it is clear that the reduction in bias is large when c is large, meaning thereby that for small sample size the proposed estimator te^{**} is superior to the estimator te^* .

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